# Space-Time Covariance functions for Planet Earth

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UTFSM at Valparaiso

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### Valparaiso



# Valparaiso



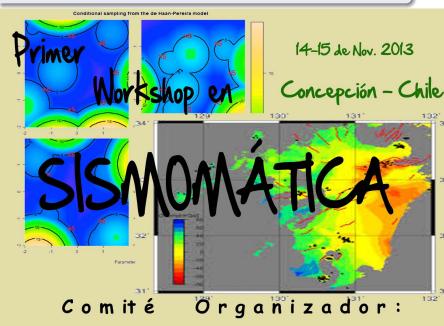
## Valparaiso







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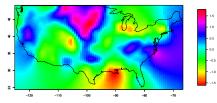




# Global Data: Kriging



Simple Kriging



# Space-time Gaussian Fields

- Assume {Z(s), s ∈ D ⊂ ℝ<sup>d</sup>}, is a Gaussian random field (GRF).
- Covariance function

$$Cov(Z(\mathbf{s}_1), Z(\mathbf{s}_2))$$

Geostatistical model

$$Y(\mathbf{s}) = \mu(\mathbf{s}) + Z(\mathbf{s})$$
  
data = deterministic part + random part

# Stationarity and Isotropy

• Weak Stationarity

 $E(Z(\mathbf{s})) = 0$   $\operatorname{Cov}(Z(\mathbf{s}), Z(\mathbf{s}')) = C(\mathbf{s} - \mathbf{s}') = C(\mathbf{h}, u)$ (h = s - s', spatial lag).

• Isotropy (Daley and Porcu, PAMS, 2013)

$$C(\mathbf{h}) = K(||\mathbf{h}||)$$

- We suppose that K comes from some parametric family  $\mathcal{C}_{ heta}$
- Example

$$C(\mathbf{h}) = K(||\mathbf{h}||; \theta) = \sigma^2 \exp(-||\mathbf{h}||/\alpha)$$

# Probabilistic Intermezzo

 Let Z = (Z(s<sub>1</sub>), Z(s<sub>2</sub>), ..., Z(s<sub>n</sub>))<sup>T</sup> the sample data. Assume it is a (partial) realization of a space-time Gaussian field Z(s).

### $\textbf{Z} \sim \textit{N}(\textbf{0}, \boldsymbol{\Sigma})$

• 
$$\Sigma = \{K(||\mathbf{s}_i - \mathbf{s}_j||)\}_{ij=1, lk=1}^n$$
 a  $n \times n$  matrix.

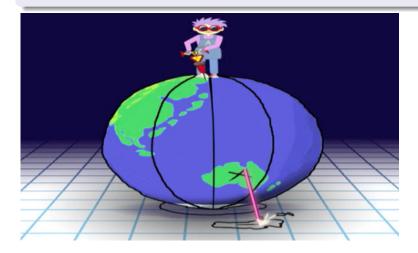
- $\Sigma \ge 0$  iff  $K(\cdot)$  is a positive definite function.
- Then practical estimation generally requires first the selection of some parametric class of positive definite functions , *i.e.*  $C(h) = C(h, \theta)$  and  $\Sigma = \Sigma(\theta)$ .

 Main goal: prediction of Z at an unknown space-time of location s<sub>0</sub> using information available from the sample data Z.

$$\hat{Z}(\mathbf{s}_0) = \mathbf{c}(\theta)^T \Sigma(\theta)^{-1} \mathbf{Z}$$
  
where  $\mathbf{c}(\theta) = \operatorname{Cov}(\mathbf{Z}, Z(\mathbf{s}_0); \theta)$ 

• We need "good" estimation of  $\theta$  to make "good" prediction.

### Spatial Stats and the Euclidean Paradigm

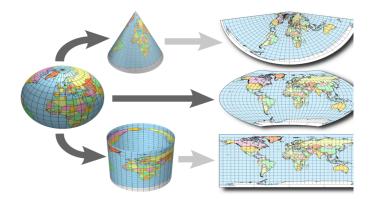


### Spatial Stats and the Euclidean Paradigm



# **Spheres**

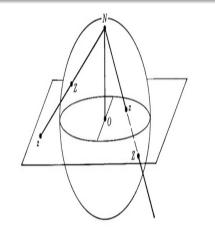
### Spatial Stats and the Euclidean Paradigm



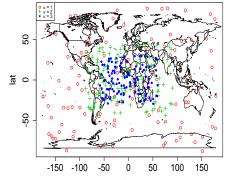
### Chordal and Great Circle

Chordal distance: criticism on

- Negative Correlations
- Counter to Spherical geometry

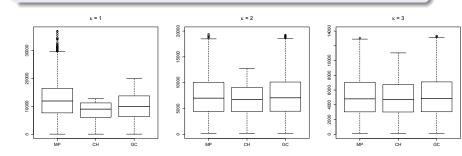


### A toy Example



lon

### A toy example (2)



#### Space-Time Challenges



Frequently, the temporal development of a process observed on a sphere is also of interest, so that the process needs to be modeled on the sphere cross time. Nevertheless, the literature on the corresponding correlation structures is sparse [...] Tilmann Gneiting, Problem 16 of Online supplement to Bernoullli (2013).

### motivations

- 2 Schoenberg coefficients and functions
- **3** The class  $\Psi_{d,T}$
- **4** Construction Principles
- Assessing discrepancies between the great circle distance and other metrics
- **6** Analysis of TOMS data

### Based on

- Porcu, E., Bevilacqua, M. & Genton, M.G. (2015). Journal of the American Statistical Association. To appear.
- **2** Berg, C. & Porcu, E. Constructive Approximation. Accepted.

#### The sphere

• d-dimensional unit sphere of  $\mathbb{R}^{d+1}$ , given as

$$\mathbb{S}^d = \{ x \in \mathbb{R}^{d+1} \mid ||x|| = 1 \}, \ d \ge 1.$$

• Great Circle distance:  $heta: \mathbb{S}^d imes \mathbb{S}^d o [0,\pi]$ ,

 $\theta(\xi,\eta) = \arccos(\xi \cdot \eta),$ 

• Chordal Distance:  $ch(\xi,\eta) = 2 \sin\left(\frac{\theta}{2}\right)$ .

#### Gaussian fields on $\mathbb{S}^d \times \mathbb{R}$

- Stationary Gaussian fields  $\{Z(\xi,t),\; (\xi,t)\in \mathbb{S}^d imes \mathbb{R}\}$ ,
- Covariance functions  $C:\mathbb{S}^d imes\mathbb{S}^d imes\mathbb{R}$ , so that

$$\operatorname{cov}\left(Z(\xi,t),Z(\eta,t')
ight)=C( heta(\xi,\eta),t-t')\,,\qquad (\xi,t),(\eta,t')\in\mathbb{S}^d imes\mathbb{R}.$$

• Class  $\Psi_{d,T}$  of continuous functions  $f: [-1,1] \times \mathbb{R} \to \mathbb{C}$  such that C can be written as

$$C(\theta(\xi,\eta),t-t') = \left[ f(\cos\theta(\xi,\eta),t-t') \right], \quad \xi,\eta \in \mathbb{S}^d, \ t,t' \in \mathbb{R}.$$
(1)

## **Spheres**

#### Schoenberg's Class: an Intermezzo

#### Iso Schoenberg

#### 21-04-1903 in Galati, 21-04-1990

- Son of a medical doctor
- 1922: M.A. at Jessy University
- 1922: Goettingen (Schur)
- 1925: Edmund Landau

#### 1930: Harvard and Princeton

- the isometric imbedding of metric spaces into Hilbert space and positive definite functions.
- 1950: Pólya
- 1966: University of Pennsylvania



#### He was...[by Richard Askey]

a man of broad culture, fluent in several languages, addicted to art, music and world literature, sensitive, gracious and giving in all ways. The working desk at his home where he engages in research is actually a drafimmn's bench complete with T-square, etc. and a tall stool. Mobiles, artistic works, models of ruled surfaces, icosahedrons and other objects are stream throughout the room. English, French and artefacts are scattered on all the nearby easy chairs.

#### The Class $\Psi_d$ : How to build it

• Gegenbauer polynomials

$$(1-2xr+r^2)^{-\lambda} = \sum_{n=0}^{\infty} C_n^{(\lambda)}(x)r^n, \quad |r| < 1, x \in \mathbb{C}.$$

• For  $\lambda > 0$ ,

$$\int_{-1}^{1} (1-x^2)^{\lambda-1/2} C_n^{(\lambda)}(x) C_m^{(\lambda)}(x) \, \mathrm{d}x = \frac{\pi \Gamma(n+2\lambda) 2^{1-2\lambda}}{\Gamma^2(\lambda)(n+\lambda) n!} \delta_{m,n}.$$

# Spheres: how to build the Class $\Psi_d$

#### The Class $\Psi_d$

• Important!

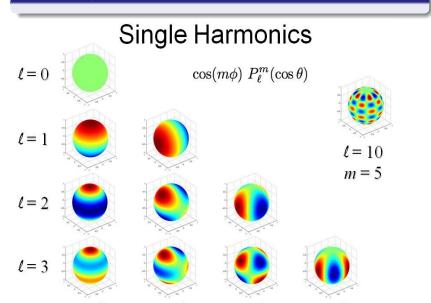
$$|C_n^{(\lambda)}(x)| \le C_n^{(\lambda)}(1), \quad x \in [-1,1].$$

- $\lambda = (d-1)/2$  and its connection with spherical harmonics.
- $n\in\mathbb{N}$ . Vector Space  $\mathcal{H}_n(d)\subset\mathcal{C}(\mathbb{S}^d)$ , dimension

 $N_n(d) := \dim \mathcal{H}_n(d) = \frac{(d)_{n-1}}{n!}(2n+d-1), \ n \ge 1, \quad N_0(d) = 1,$ 

# Spheres: how to build the Class $\Psi_d$

### The Class $\Psi_{d,T}$



# Spheres: how to build the Class $\Psi_d$

#### The Class $\Psi_d$

• Orthogonality relation:

$$\int_{-1}^{1} (1-x^2)^{d/2-1} c_n(d,x) c_m(d,x) \, \mathrm{d}x = \frac{||\omega_d||}{||\omega_{d-1}||N_n(d)} \delta_{m,n}.$$

• Theorem(Schoenberg, 1942) A continuous function  $f: [-1,1] \to \mathbb{R}$  belongs to the class  $\Psi_d$ ,  $d = 1, 2, \ldots$ , if and only if

$$f(\cos \theta) = \sum_{n=0}^{\infty} b_{n,d} c_n(d, \cos \theta), \quad b_{n,d} \ge 0, \theta \in [0, \pi].$$

for a summable sequence  $(b_{n,d})_{n=0}^\infty$  given as

$$b_{n,d} = \frac{||\omega_{d-1}||N_n(d)}{||\omega_d||} \int_{-1}^1 f(x)c_n(d,x)(1-x^2)^{d/2-1} \,\mathrm{d}x.$$

#### The Class $\Psi_{d,T}$

- If f(0) = 1,  $(b_{n,d})$  is a probability sequence
- Daley and Porcu (2013) call  $b_{n,d}$  d-Schoenberg coefficients and the sequence  $(b_{n,d})$  a d-Schoenberg sequence

• In 
$$d = 1$$
,

$$f(\cos \theta) = \sum_{n=0}^{\infty} b_{n,1} \cos(n\theta), \quad b_{n,1} \ge 0, \theta \in [0,\pi],$$

#### Characterization of $\Psi_{d,T}$

$$\begin{split} & a_1 \to a_k(r), \quad \frac{a_1r}{a_k} \;, \quad a_k(r) \to a(k) \\ & a_k(r) \in \tilde{e}^{kr} \quad \int_{a_1}^{a_2r} a_k(r,\tilde{r}) + e^{ikr} \quad \tilde{T}_k \to \left[ \Psi \left[ p_1 \pi \right] \to \mathcal{R} : \quad H\Psi \right] \\ & \left[ \frac{a_2r}{a_2} \frac{d_2}{d_2} \right] \end{split}$$
 $\begin{array}{c} & \varphi \in \left( \left\| u^{(1)}, \varphi(u, \cdot), \varphi(u, \cdot) \right\| \leq \ell \cdot u \\ \varphi(u, u^{(1)}, - z_0) \quad \forall \ \varphi \in \mathbb{R}^d \\ \varphi(u, u^{(1)}, - z_0) \quad \forall \ \varphi \in \mathbb{R}^d \\ \varphi(u, u^{(1)}, u^{(1)}, u^{(1)}) \\ \varphi(u, u^{(1)}, u^{(1)})$  $\int_{0}^{1} \int_{0}^{1} \frac{d^{2}}{dt} \int_{0}^{1} \int_{0}^{1} \frac{d^{2}}{dt} \int_{0}^{1} \frac{d^{2}}{dt}$ und and unlard means with complexity sin for the time the complexity of the transformed and the second seco ( of any pro MEE

#### The Class $\Psi_{d,T}$ : Characterization Theorems

**Theorem 1** Let  $d \in \mathbb{N}$  and let  $f: [-1,1] \times \mathbb{R} \to \mathbb{C}$  be a continuous function. Then f belongs to  $\Psi_{d,T}$  if and only if there exists a sequence  $\varphi_{n,d}: \mathbb{R} \to \mathbb{R}$  of p.d. functions with  $\sum \varphi_{n,d}(0) < \infty$  such that

$$f(\cos\theta,t) = \sum_{n=0}^{\infty} \varphi_{n,d}(t) c_n(d,\cos\theta),$$

and the above expansion is uniformly convergent for  $( heta,t)\in [0,\pi] imes \mathbb{R}$ . We have

$$\varphi_{n,d}(t) = \frac{N_n(d)||\omega_{d-1}||}{||\omega_d||} \int_{-1}^1 f(x,t) c_n(d,x) (1-x^2)^{d/2-1} \,\mathrm{d}x.$$

# Construction principles

### with Moreno Bevilacqua and Marc Genton



#### The Class $\Psi_{d,T}$

A natural construction:

$$Z(\eta,t) = \sum_{k=0}^{\infty} \sum_{\nu \in \Upsilon_{k,d}} \xi_{k,\nu}(t) Y_{k,\nu,d}(\eta), \qquad \eta \in \mathbb{S}^d, t \in \mathbb{R},$$

- $Y_{k,\nu,d}:\mathbb{S}^d o \mathbb{C}$ : normalized hyperspherical harmonics;
- $\xi_{k,\nu}(t)$ : Gaussian processes, with zero mean and  $\mathbb{E}\xi_{k,\nu}(t)\xi_{k',\nu'}(t') = \delta_{k,k'}\delta_{\nu,\nu'}g_k(t-t')$ ,  $t,t' \in \mathbb{R}$ .

#### An easy construction principle

Let  $\{g_k(\cdot)\}_{k=0}^{\infty}$  be an absolutely convergent sequence of continuous and positive definite functions on the real line, such that  $g_k(0) = b_k$  for all  $k = 0, 1, \ldots$ , with  $\{b_k\}_{k=0}^{\infty}$  being a probability mass sequence. Then,

$$C( heta, u) = \sum_{k=0}^{\infty} g_k(u) (\cos heta)^k, \qquad ( heta, u) \in [0, \pi] imes \mathbb{R},$$

is a representation for members of the class  $\Psi_{\infty,\mathcal{T}}$ .

#### The Class $\Psi_{d,T}$

| Family                  | Analytic expression  | Parameters range  |
|-------------------------|--|---|
| Negative Binomial       | $C(\theta, u) = \left\{ \frac{1 - \varepsilon}{1 - \varepsilon g(u) \cos \theta} \right\}^{T}$                       | $arepsilon\in(0,1),	au>0$   |
| Multiquadric            | $C(\theta, u) = \frac{(1-\varepsilon)^{2\tau}}{\left\{1+\varepsilon^2 - 2\varepsilon g(u)\cos\theta\right\}^{\tau}}$ | $arepsilon\in(0,1),\ 	au>0$   |
| Sine Series             | $C(\theta, u) = e^{g(u) \cos \theta - 1} \left\{ 1 + g(u) \cos \theta \right\} / 2$                                  |   |
| Sine Power              | $C(\theta, u) = 1 - 2^{-\alpha} \{1 - g(u) \cos \theta\}^{\alpha/2}$   | $lpha\in (0,2]$   |
| Adapted<br>Multiquadric | $C(\theta, u) = \left[\frac{\{1+g^{2}(u)\}(1-\varepsilon)}{1+g^{2}(u)-2\varepsilon g(u)\cos\theta}\right]^{\tau}$    | $arepsilon \in (0,1), 	au > 0$<br>$2g(\cdot)/\{1+g^2(\cdot)\}$<br>corr. function on $\mathbb R$ |
| Poisson                 | $C(	heta, u) = \exp\left[\lambda \left\{\cos 	heta g(u) - 1 ight\} ight]$  | $\lambda > 0$   |

### Spatial Adapting from the Gneiting class

• Gaussian process Z on  $\mathbb{R}^d \times \mathbb{R}$ , points  $(\mathbf{x}, t)$  and  $(\mathbf{y}, t')$  such that  $\|\mathbf{y} - \mathbf{x}\| = \mathbf{h}$  (with  $\|\cdot\|$  denoting the Euclidean distance) and t - t' = u,

$$C(\boldsymbol{h}, \boldsymbol{u}) := \frac{\sigma^2}{\psi(\|\boldsymbol{h}\|^2)^{d/2}} \varphi\left\{\frac{\boldsymbol{u}^2}{\psi(\|\boldsymbol{h}\|^2)}\right\}, \qquad (\boldsymbol{h}, \boldsymbol{u}) \in \mathbb{R}^d \times \mathbb{R},$$

•  $\varphi$  is completely monotone on the positive real line such that  $\varphi(0) = 1$ ,  $\psi$  is a positive-valued Bernstein function, and  $\sigma^2$  is a variance parameter.

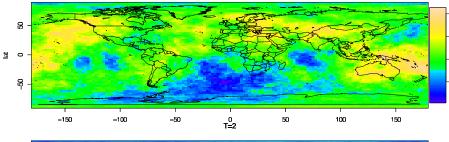
### Spatial Adapting from the Gneiting class

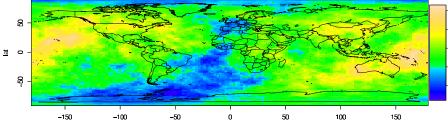
Let  $\theta: \mathbb{S}^d \times \mathbb{S}^d \to [0,\pi]$  be the great circle distance. Let  $\varphi: [0,\infty) \to \mathbb{R}_+$  be a completely monotone function on the positive real line, with  $\varphi(0) = 1$ , and let  $\psi$  be a positive-valued Bernstein function. Denote by  $\psi_{[0,\pi]}$  the restriction of  $\psi$  to the interval  $[0,\pi]$ . Then, the function

$$C(\theta, u) := \frac{\sigma^2}{\psi_{[0,\pi]}(\theta)^{1/2}} \varphi\left\{\frac{u^2}{\psi_{[0,\pi]}(\theta)}\right\}, \qquad (\theta, u) \in [0,\pi] \times \mathbb{R}, \quad (2)$$

belongs to the class  $\Psi_{\infty,\mathcal{T}}$ .

## Gneiting class on the sphere





T=1

## **Relevant Comments**

- Mean square differentiability for processes on spheres. Attempts in Jeong and Jun (2015).
- Exception being the Sine Power model.
- Adapted construction has the same problem: Matérn is valid only for  $u \in (0, 1/2].$
- Adapted construction allows for rescaling the spatial component, direct construction not
- Direct construction allow for any type of temporal margin, provided g is a temporal correlation function.

## Examples from The Adapted Gneiting Class

$$C(\theta, u) = \frac{\sigma^2}{\left\{1 + \left(\frac{R\theta}{c_s}\right)^{\alpha}\right\}^{\delta + \beta/2}} \exp\left[-\frac{\left(\frac{|u|}{c_\tau}\right)^{2\gamma}}{\left\{1 + \left(\frac{R\theta}{c_s}\right)^{\alpha}\right\}^{\beta\gamma}}\right],$$

$$C(\theta, u) = \frac{\sigma^2}{\left\{1 + \left(\frac{R\theta}{c_s}\right)^{\alpha}\right\}^{\delta + \beta/2}} \left[1 + \frac{\left(\frac{|u|}{c_T}\right)^{2\tau}}{\left\{1 + \left(\frac{R\theta}{c_s}\right)^{\alpha}\right\}^{\tau\beta}}\right]^{-\lambda},$$

### Examples from The Adapted Gneiting Class

• Take the negative binomial family and  $u \mapsto g(u; \alpha) := (1 + |u|^{\alpha})^{-1}$ ,  $\alpha \in (0, 2]$ ,

$$C(\theta, u) = \sigma^{2} \left[ \frac{1 - \varepsilon}{1 - \varepsilon \left\{ 1 + \left( \frac{|u|}{c_{T}} \right)^{\alpha} \right\}^{-1} \cos \theta} \right]^{\tau}, \qquad (\theta, u) \in [0, \pi] \times \mathbb{R},$$

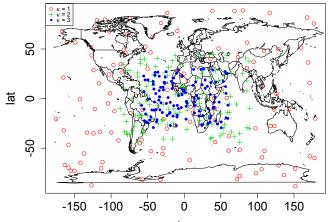
From the multiquadric,

$$C(\theta, u) = \frac{\sigma^{2}(1-\varepsilon)^{2\tau}}{\left[1+\varepsilon^{2}-2\varepsilon\left\{1+\left(\frac{|u|}{c_{T}}\right)^{\alpha}\right\}^{-1}\cos\theta\right]^{\tau}}, \qquad (\theta, u) \in [0, \pi] \times \mathbb{R},$$
(3)

with the same restriction on the parameters as in the previous model.

$$C(\theta, u) = \sigma^{2} \upsilon \left[ 1 + \frac{\cos \theta}{c_{S} \left\{ 1 + \left(\frac{|u|}{c_{T}}\right)^{\alpha} \right\}^{-1}} \right] \exp \left[ \frac{\cos \theta}{c_{S} \left\{ 1 + \left(\frac{|u|}{c_{T}}\right)^{\alpha} \right\}^{-1}} \right], \qquad (\theta, u) \in [0, \pi] \times \mathbb{R},$$

## **Simulation Studies**



- Three scenarios;
- Estimate  $\lambda$  using ML;

• Under 
$$C(\theta, u; \lambda) = \frac{\sigma^2}{\left\{1 + \left(\frac{R\theta}{c_S}\right)\right\}} \exp\left[-\frac{|u|}{c_T \left\{1 + \left(\frac{R\theta}{c_S}\right)\right\}^{1/4}}\right],$$

- Using either the GC, CH or MP distances. Notation  $\widehat{\lambda}^{(k)}_{\mathcal{X}}$ , with  $\mathcal{X}=$  GC, CH or MP
- $k = 1, \ldots, 1000$  Simulations.

### Scenarios

Discrepancy between the ML estimates using either GC, CH and MP distances. Given  $\widehat{\lambda}^{(k)}_{\mathcal{X}}$ , we call  $M(\cdot)$  the measure

$$M^{\mathcal{X}}(i\widehat{\lambda}) = \sqrt{\frac{\sum_{k=1}^{1000} (i\widehat{\lambda}_{\mathrm{GC}}^{(k)} - i\widehat{\lambda}_{\mathcal{X}}^{(k)})^2}{1000}}, \qquad i = 1, 2, 3, \qquad \mathcal{X} = \mathrm{CH}, \mathrm{MP}.$$
(4)

We also define another measure  $A(\cdot)$  by

$$\mathcal{A}^{\mathcal{X}}(i\widehat{\lambda}) = \sqrt{\frac{\sum_{k=1}^{1000} (i\widehat{\lambda}_{\mathcal{X}}^{(k)} - i\lambda)^2}{1000}}, \qquad i = 1, 2, 3, \qquad \mathcal{X} = \text{GC, CH, MP},$$
(5)

where  $_i\lambda$  denotes the nominal value chosen under one of the proposed scenarios.

## **Simulation Studies**

|                          |        | $\kappa = 1$ |       | $\kappa = 2$ |       | $\kappa = 3$ |
|--------------------------|--------|--------------|-------|--------------|-------|--------------|
|                          | (I)    | (II)         | (I)   | (II)         | (I)   | (II)         |
| $M^{CH}(\hat{c}_S)$      | 6.25   | 24.07        | 0.97  | 4.06         | 0.42  | 2.19         |
| $M^{MP}(\hat{c}_S)$      | 93.03  | 201.99       | 9.86  | 17.45        | 3.17  | 5.91         |
| $M^{CH}(\hat{c}_T)$      | 0.013  | 0.014        | 0.004 | 0.008        | 0.004 | 0.007        |
| $M^{MP}(\hat{c}_T)$      | 0.058  | 0.080        | 0.015 | 0.022        | 0.007 | 0.012        |
| $M^{CH}(\hat{\sigma}^2)$ | 0.004  | 0.011        | 0.001 | 0.003        | 0.001 | 0.002        |
| $M^{MP}(\hat{\sigma}^2)$ | 0.026  | 0.045        | 0.005 | 0.008        | 0.002 | 0.004        |
| $A^{GC}(\hat{c}_S)$      | 116.47 | 206.77       | 67.74 | 111.52       | 45.67 | 74.90        |
| $A^{CH}(\hat{c}_{S})$    | 117.73 | 211.78       | 68.00 | 112.34       | 45.73 | 74.98        |
|                          |        |              |       |              |       |              |
| $A^{MP}(\hat{c}_S)$      | 133.73 | 261.39       | 68.59 | 112.96       | 45.89 | 74.16        |
| $A^{GC}(\hat{c}_T)$      | 0.212  | 0.211        | 0.212 | 0.212        | 0.212 | 0.211        |
| $A^{CH}(\hat{c}_T)$      | 0.212  | 0.212        | 0.212 | 0.212        | 0.212 | 0.212        |
| $A^{MP}(\hat{c}_T)$      | 0.220  | 0.226        | 0.213 | 0.213        | 0.213 | 0.212        |
| $A^{GC}(\hat{\sigma}^2)$ | 0.088  | 0.101        | 0.084 | 0.094        | 0.083 | 0.093        |
| $A^{CH}(\hat{\sigma}^2)$ | 0.088  | 0.104        | 0.084 | 0.094        | 0.083 | 0.093        |
| $A^{MP}(\hat{\sigma}^2)$ | 0.089  | 0.105        | 0.084 | 0.095        | 0.083 | 0.093        |

- Level-3 Total Ozone Mapping Spectrometer (TOMS): daily total column ozone levels.
- Spatially irregular grid (1° latitude by 1.25° longitude away from the poles)
- Original data: Latitude interval [-89.5,89.5] and longitudes [-180,180]
- Jun and Stein (2008): spatial dataset
- Here: 15 obs. in time, for a total of of 20,160 points (288 longitudinal and 70 latitudinal) observed during 15 days, for a total of 302,400 observations.

- For the missing data: follow Jun and Stein (2008): local averaging (24 observations) for each local averaging.
- Likelihood estimation unfeasible: select a subgrid of 336 spatial points and all temporal observations, for a total of 5,040 observations.
- Detrend the data using spatio-temporal splines
- Residuals as a realization from a zero mean space-time Gaussian random field.

- A. Two models based on the adapted Gneiting classes
- B. Three models based on direct construction, hence valid with GC only.
- C. A model based on the Gneiting class valid using CH and MP distances

| Distance       | GC       | CH       | MP       | GC       | CH       | MP       | GC       | CH       | MP       |
|----------------|----------|----------|----------|----------|----------|----------|----------|----------|----------|
| Model          |          | A.1      |          |          | A.2      |          |          | C.1      |          |
| c <sub>S</sub> | 742.7    | 743.9.3  | 734.2    | 681.3    | 672.2    | 733.3    | -        | 450.8    | 417.5    |
| c <sub>T</sub> | 2.54     | 2.54     | 2.17     | 1.76     | 1.67     | 1.71     | -        | 1.56     | 1.29     |
| $\beta^{c_T}$  | 1        | 1        | 0.95     | 1        | 1        | 1        | -        | 0.95     | 0.89     |
| $\sigma^2$     | 102.6    | 102.1    | 102.7    | 106.1    | 103.0    | 111.4    | -        | 97.9     | 98.5     |
| Nugget         | 9.81     | 9.82     | 6.35     | 6.30     | 5.60     | 4.97     | -        | 12.26    | 5.72     |
| Likelihood     | -17233.8 | -17234.0 | -17296.5 | -17257.1 | -17258.5 | -17317.3 | -        | -17156.3 | -17223.9 |
| Model          |          | B.1      |          |          | B.2      |          |          | B.3      |          |
| $\tau$         | 0.01     | -        | -        | 144.09   | -        | -        | 4.04     | -        | -        |
| CT CT          | 51.05    | -        | -        | 50.58    | -        | -        | 36.85    | -        | -        |
| α              | 1.49     | -        | -        | 1.49     | -        | -        | 1.61     | -        | -        |
| $\sigma^2$     | 85.8     | -        | -        | 86.0     | -        | -        | 89.9     | -        | -        |
| Nugget         | 18.60    | -        | -        | 18.54    | -        | -        | 16.84    | -        | -        |
| Likelihood     | -17168.3 | -        | -        | -17167.9 | -        | -        | -17162.8 | -        | -        |

## Caesar: Ad memoriam

Caesar: 28th july 2014–15th June 2014

